Univariate and Multivariate Analysis of Time Series and Complex Networks

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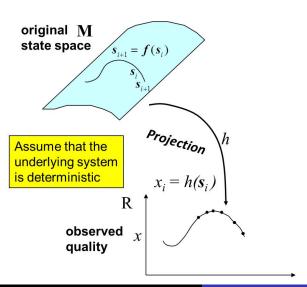


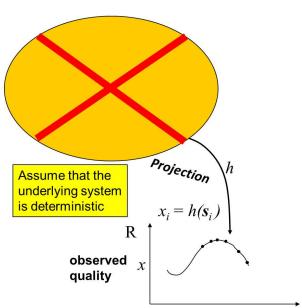
Dependence measures in univariate time series

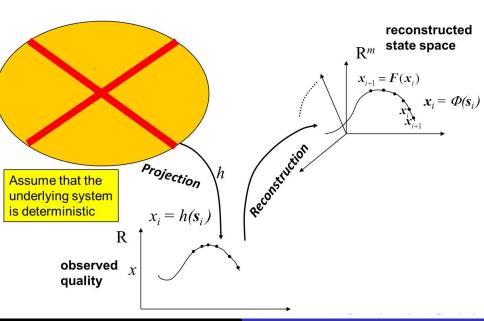
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- Interdependence in multivariate time series

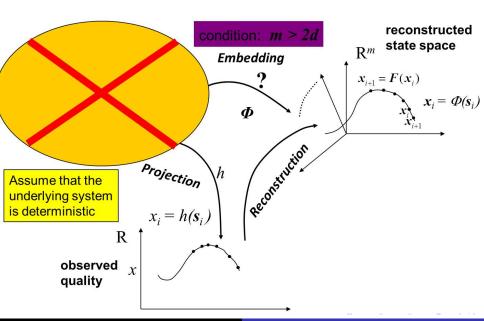
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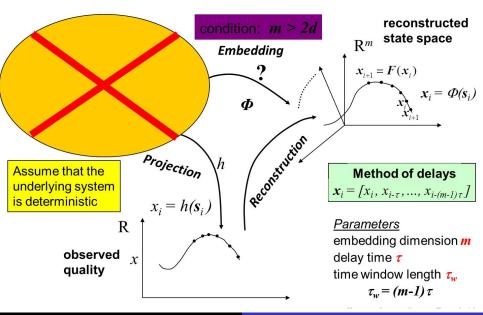
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- 3 Complex networks from multivariate time series
- High-dimensional time series: Implications and solutions











... Y_{t-5} Y_{t-4} Y_{t-3} Y_{t-2} Y_{t-1} Y_{t} Y_{t+1}

t-5 t-4 t-3 t-2 t-1 t t+1

Dynamical system

$$s_{t+1} = f(s_t), \quad s_t \in \mathbb{R}^d$$

 $\dot{s} = f(s), \quad s_t \in \mathbb{R}^d$

$$\dots \quad Y_{t\text{-}5} \qquad Y_{t\text{-}4} \qquad Y_{t\text{-}3} \qquad Y_{t\text{-}2} \qquad Y_{t\text{-}1}$$

$$Y_{t-1}$$

$$Y_{t-1}$$
 Y_t

$$Y_{t+1}$$

Observable (assuming $y_t = h(s_t)$)

$$y_{t+1} = F(\mathbf{y}_t) = F(y_t, y_{t-1}, ..., y_{t-m+1}) \quad y_t \in \mathbb{R}$$

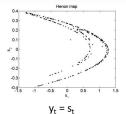
t-5 t-4 t-3

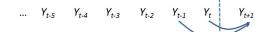
t-2

t-1

t+1

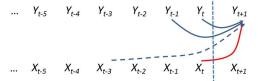
$$\begin{split} s(t) &= 1 - 1.4 \, s(t\text{-}1)^2 \, + \, 0.3 s(t\text{-}2) \\ or \\ s_1(t) &= 1 - 1.4 \, s_1(t\text{-}1)^2 \, + \, s_2(t\text{-}1) \\ s_2(t) &= 0.3 \, s_1(t\text{-}1) \end{split}$$



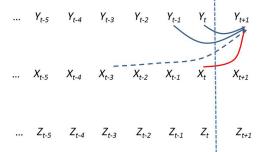


· t-5 t-4 t-3 t-2 t-1 t t+1

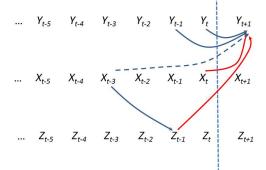


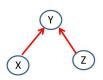


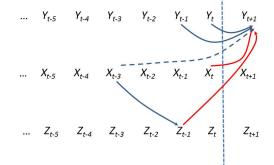
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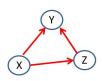
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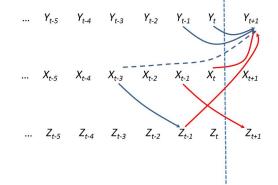






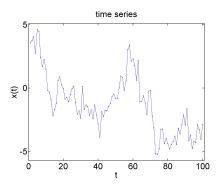
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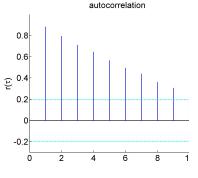
Are X_t and X_{t-1} linearly correlated? Are X_t and X_{t-2} linearly correlated?

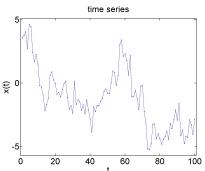


Autocorrelation $r(\tau) = r(X_t; X_{t-\tau})$

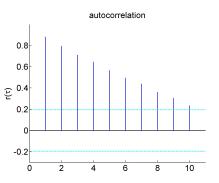
Are X_t and X_{t-1} linearly correlated? $r(1) = r(X_t; X_{t-1}) \neq 0$? Yes Are X_t and X_{t-2} linearly correlated? $r(2) = r(X_t; X_{t-2}) \neq 0$? Yes

$$r(\tau) = r(X_t; X_{t-\tau}) = \frac{1}{n-\tau} \sum_{t=\tau+1}^{n} (x_t - \bar{x})(x_{t-\tau} - \bar{x})/s_X^2$$





Are X_t and X_{t-2} directly linearly correlated?

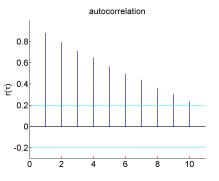


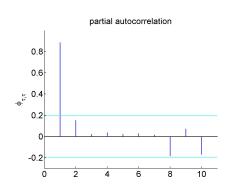
Partial autocorrelation $\phi_{\tau,\tau} = r(X_t; X_{t-\tau}|X_{t-1}, \dots, X_{t-\tau+1})$

Are X_t and X_{t-2} directly linearly correlated?

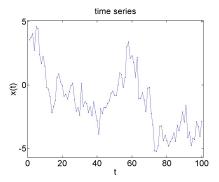
Are X_t and X_{t-2} linearly correlated given X_{t-1} ?

$$r(X_t; X_{t-2}|X_{t-1}) \neq 0$$
? No





Are X_t and X_{t-1} linearly and nonlinearly correlated? Are X_t and X_{t-2} linearly and nonlinearly correlated?

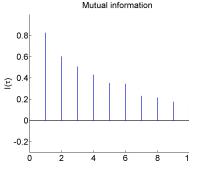


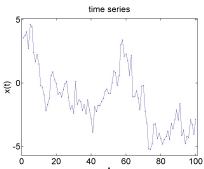
Mutual information $I(\tau) = I(X_t; X_{t-\tau})$

Are X_t and X_{t-1} linearly and nonlinearly correlated? $I(X_t; X_{t-1}) \neq 0$? Yes Are X_t and X_{t-2} linearly and nonlinearly correlated? $I(X_t; X_{t-2}) \neq 0$? Yes $I(\tau) = I(X_t, X_{t-\tau}) = H(X_t) + H(X_{t-\tau}) - H(X_t, X_{t-\tau})$

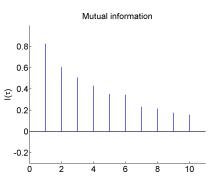
$$I(\tau) = I(X_t, X_{t-\tau}) = H(X_t) + H(X_{t-\tau}) - H(X_t, X_{t-\tau})$$

$$= \sum_{X_t, X_{t-\tau}} p_{X_t X_{t-\tau}}(x_t, x_{t-\tau}) \log \frac{p_{X_t X_{t-\tau}}(x_t, x_{t-\tau})}{p_{X_t}(x_t) p_{X_{t-\tau}}(x_{t-\tau})}$$





Are X_t and X_{t-2} directly linearly and nonlinearly correlated?

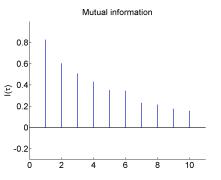


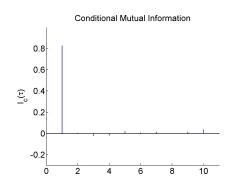
Conditional mutual information $I_c(\tau) = I(X_t; X_{t-\tau}|X_{t-1}, \dots, X_{t-\tau+1})$

Are X_t and X_{t-2} directly linearly and nonlinearly correlated?

Are X_t and X_{t-2} linearly and nonlinearly correlated given X_{t-1} ?

$$I(X_t; X_{t-2}|X_{t-1}) \neq 0$$
? No



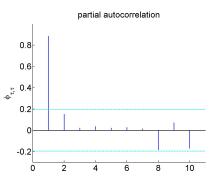


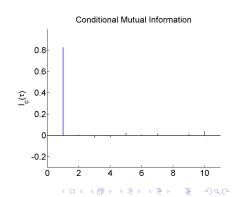
Are X_t and X_{t-2} directly linearly correlated?

$$r(X_t; X_{t-2}|X_{t-1}) \neq 0$$
? No

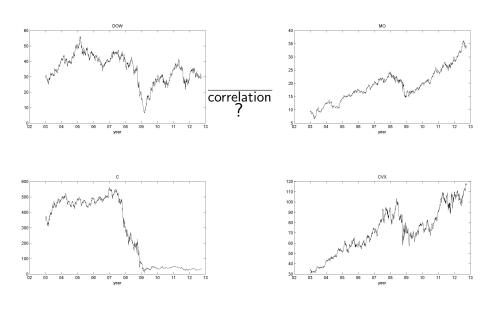
Are X_t and X_{t-2} directly linearly or/and nonlinearly correlated?

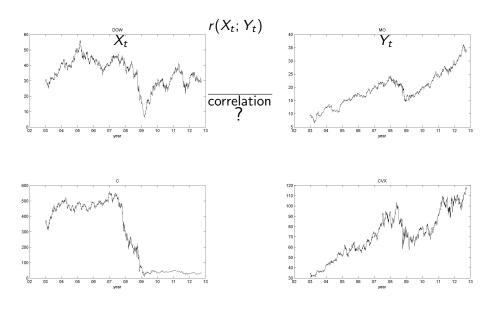
$$I(X_t; X_{t-2}|X_{t-1}) \neq 0$$
? No

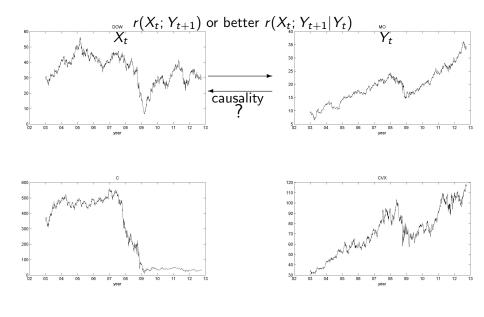


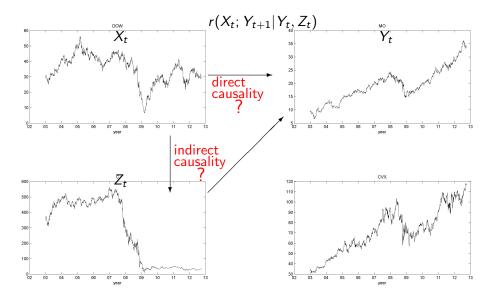


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Correlation measures

Bivariate time series $\{x_t, y_t\}_{t=1}^n$

Linear correlation measures:

Estimate of cross-covariance

$$c_{XY}(\tau) = \hat{\gamma}_{XY}(\tau) = \frac{1}{n-\tau} \sum_{t=1}^{n-\tau} (x_t - \bar{x})(y_{t+\tau} - \bar{y})$$

 \bar{x} and \bar{y} are sample means.

Estimate of cross-correlation:

$$r_{XY}(\tau) = \hat{\rho}_{XY}(\tau) = \frac{c_{XY}(\tau)}{c_{XY}(0)} = \frac{c_{XY}(\tau)}{s_X s_Y}$$

 s_X and s_Y are sample standard deviations.

- $|r_{XY}(\tau)| \le 1$
- $r_{XY}(\tau) = r_{YX}(-\tau)$ but $r_{XY}(\tau) \neq r_{XY}(-\tau)$



Nonlinear correlation measures:

Entropy: information from each sample of X (assume proper discretization of X)

$$H(X) = -\sum_{x} p_X(x) \log p_X(x)$$

Mutual information: information for Y knowing X and vice versa

$$I(X, Y) = H(X) + H(Y) - H(X, Y) = \sum_{x,y} p_{XY}(x, y) \log \frac{p_{XY}(x, y)}{p_X(x)p_Y(y)}$$

For $X \to X_t$ and $Y \to Y_{t+\tau}$, cross-delayed mutual information:

$$I_{XY}(\tau) = I(X_t, Y_{t+\tau}) = \sum_{x_t, y_{t+\tau}} p_{X_t Y_{t+\tau}}(x_t, y_{t+\tau}) \log \frac{p_{X_t Y_{t+\tau}}(x_t, y_{t+\tau})}{p_{X_t}(x_t) p_{Y_{t+\tau}}(y_{t+\tau})}$$

To compute $I_{XY}(\tau)$ make a partition of $\{x_t\}_{t=1}^n$, a partition of $\{y_t\}_{t=1}^n$ and compute probabilities for each cell from the relative frequency.

$$r_{XY}(0) \neq 0$$
:

- \implies (linear) correlation of x_t and y_t
- \implies systems X and Y are correlated, $X \sim Y$

$$r_{XY}(\tau) \neq 0$$
:

- \implies (linear) correlation of x_t and $y_{t+\tau}$
- $\Longrightarrow X$ effects the future of Y

$$\Longrightarrow X \to Y$$

$$r_{XY}(-\tau) \neq 0 \implies Y \rightarrow X$$

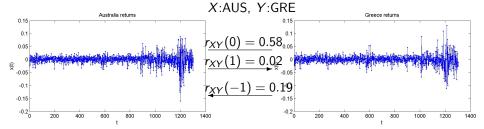
Thus $r_{XY}(\tau)$ and $I_{XY}(\tau)$ indicate the direction of interaction.

Can they also be used as causality measures?

Not the most appropriate, but they have been used in many studies

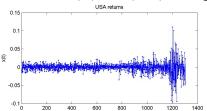


Example: Returns for USA, UnitedKingdom, Greece and Australia.

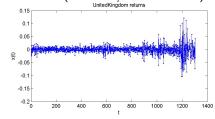


returns:

$$x_t = \log(y_t) - \log(y_{t-1})$$



Is the measure significant? Can I draw a link? (directed/non-directed)



Significance randomization test for a correlation / causality measure q, $H_0: q = 0$ $H_1: q \neq 0$

• Generate M resampled (surrogate) time series, each by shifting the original observations with a random time step w:

original time series: $\{x_t\} = \{x_1, x_2, \dots, x_n\}$

i-th surrogate time series:

$$\{x_t^{*i}\} = \{x_{w+1}, x_{w+2}, \dots, x_n, x_1, \dots, x_{w-1}, x_w\}$$

2 Compute the statistic q on the original pair, q_0 , and on the M surrogate pairs, q_1, \ldots, q_M ,

e.g.
$$q_0 \equiv r_{XY}(\tau) = \mathsf{Corr}(x_t, y_{t+\tau})$$
 and $q_i \equiv \mathsf{Corr}(x_t^{*i}, y_{t+\tau}^{*i})$

3 If q_0 is at the tails of the empirical null distribution formed by q_1, \ldots, q_M , reject H_0 .

Using rank ordering: for a two-sided test, the p-value of the test is

$$2\frac{r_{q_0}-0.326}{M+1+0.348} \quad \text{if} \quad r_{q_0} < \frac{M+1}{2} \\ 2\left(1-\frac{r_{q_0}-0.326}{M+1+0.348}\right) \quad \text{if} \quad r_{q_0} \geq \frac{M+1}{2} \\ \frac{1}{2} \\ \frac{1}{2$$

Example: Returns for USA, UnitedKingdom, Greece and Australia. Correlation matrix for delay 1, $r_{XY}(1)$

$$R(1) = \begin{bmatrix} 0.382 & 0.333 & 0.596 \\ 0.049 & 0.039 & 0.303 \\ 0.096 & 0.001 & 0.190 \\ 0.031 & -0.001 & -0.021 \end{bmatrix}$$

Randomization significance test for $r_{XY}(1)$ (M = 1000)

Matrix of p-values

Adjacency matrix

$$P(R(1)) = \begin{bmatrix} 0.0013 & 0.0013 & 0.0033 \\ 0.0732 & 0.1991 & 0.0013 \\ 0.0073 & 0.8901 & 0.0033 \\ 0.2450 & 0.9760 & 0.4028 \end{bmatrix} A = \begin{bmatrix} 1 & 1 & 1 \\ 0 & 0 & 1 \\ 1 & 0 & 1 \\ 0 & 0 & 0 \end{bmatrix}$$

For significance level, say $\alpha=0.05$, there may be $p<\alpha$ more often than it should be due to multiple testing. Correction with e.g. False Discovery Rate (FDR)



Linear causality measures (direct and indirect)

Idea of Granger causality $X \to Y$ [Granger 1969]: predict Y better when including X in the regression model.

Granger Causality Index (GCI) [Brandt & Williams 2007]

Bivariate time series $\{x_t, y_t\}_{t=1}^n$ driving system: X, response system: Y

Model 1 (restricted, R, X absent in the model):

$$y_t = \sum_{i=1}^p a_i y_{t-i} + e_{R,t}$$

Model 2 (unrestricted, U, X present in the model):

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + e_{U,t}$$

$$\mathsf{GCI}_{X \to Y} = \mathsf{In} \, \frac{\mathsf{Var}(\hat{e}_{R,t})}{\mathsf{Var}(\hat{e}_{U,t})} \qquad \mathsf{GCI}_{X \to Y} > 0 \Rightarrow X \to Y \; \mathsf{holds}$$

Parametric significance test for GCI

$$GCI_{X\to Y} > 0$$
 ? \Rightarrow Significance test

If X does not Granger causes Y then the contribution of X-lags in the unrestricted model should be insignificant \Rightarrow

the terms of X should be insignificant

$$\mathsf{H}_0$$
: $b_i = 0$, for all $i = 1, \ldots, p$
 H_1 : $b_i \neq 0$, for any of $i = 1, \ldots, p$

Snedecor-Fisher test (F-test):

$$F = \frac{(SSE^R - SSE^U)/p}{SSE^U/ndf}$$

SSE: sum of squared errors

ndf: number of degrees of freedoms, ndf = (n - p) - 2p,

n-p: number of equations,

2p: number of coefficients in the U-model.



Linear causality measures (direct)

Conditional Granger Causality Index (CGCI)

K time series $\{x_t, y_t\}_{t=1}^n$ and $\{\mathbf{z}_t\}_{t=1}^n = \{z_{1,t}, z_{2,t}, \dots, z_{K-2,t}\}_{t=1}^n$ driving system: X, response system: Y, conditioning on system Z, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$

Model 1 (restricted, R, X absent in the model):

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} A_i \mathbf{z}_{t-i} + e_{R,t}$$

Model 2 (unrestricted, U, X present in the model):

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + \sum_{i=1}^p A_i \mathbf{z}_{t-i} + e_{U,t}$$

$$\frac{\mathsf{CGCI}_{X \to Y|Z}}{\mathsf{Var}(\hat{e}_{U,t})} = \mathsf{In}\, \frac{\mathsf{Var}(\hat{e}_{R,t})}{\mathsf{Var}(\hat{e}_{U,t})}$$



Parametric significance test for CGCI

$$CGCI_{X \to Y|Z} > 0$$
 ? \Rightarrow Significance test as for GCI
H₀: $b_i = 0$, for all $i = 1, ..., p$
H₁: $b_i \neq 0$, for any of $i = 1, ..., p$

$$F = \frac{(SSE^R - SSE^U)/p}{SSE^U/ndf}$$

$$ndf = (n - p) - Kp$$
,
 $n - p$: number of equations,

Kp: number of coefficients in the U-model.

VAR model for Y

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + e_{U,t}$$

VAR model for Y

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

$$y_{t+1} = \sum_{i=1}^{p} a_i y_{t-i+1} + \sum_{i=1}^{p} b_i x_{t-i+1} + e_{U,t+1}$$

 y_{t+1} is given in terms of $\{y_t,y_{t-1},\ldots,y_{t-p+1}\}$ and $\{x_t,x_{t-1},\ldots,x_{t-p+1}\}.$

VAR model for Y

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + e_{U,t}$$

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 $\mathbf{y}_t = [y_t, y_{t-1}, \dots, y_{t-p+1}]$: vector of lagged Y

let the lag step be $\tau \geq 1 \quad \Rightarrow \quad \mathbf{y}_t = [y_t, y_{t-\tau}, \dots, y_{t-(p-1)\tau}]$:

au, p: embedding parameters (generally different for X and Y)

VAR model for Y

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 τ , p: embedding parameters (generally different for X and Y)

State space reconstruction:

$$\mathbf{x}_t = [x_t, x_{t-\tau_x}, \dots, x_{t-(m_x-1)\tau_x}]'$$
, embedding parameters: m_x, τ_x
 $\mathbf{y}_t = [y_t, y_{t-\tau_y}, \dots, y_{t-(m_y-1)\tau_y}]'$, embedding parameters: m_y, τ_y

 y_{t+1} : future state of Y



Nonlinear causality measures (direct and indirect)

Transfer Entropy (TE) [Schreiber, 2000]

Measure the effect of X on Y at one time step ahead, accounting (conditioning) for the effect from its own current state

$$TE_{X \to Y} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t)$$

$$= H(\mathbf{x}_t, \mathbf{y}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) + H(y_{t+1}, \mathbf{y}_t) - H(\mathbf{y}_t)$$

$$= \sum p(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) \log \frac{p(y_{t+1} | \mathbf{x}_t, \mathbf{y}_t)}{p(y_{t+1} | \mathbf{y}_t)}$$

Nonlinear causality measures (direct and indirect)

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$$= H(\mathbf{x}_t, \mathbf{y}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) + H(y_{t+1}, \mathbf{y}_t) - H(\mathbf{y}_t)$$

$$= \sum p(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) \log \frac{p(y_{t+1} | \mathbf{x}_t, \mathbf{y}_t)}{p(y_{t+1} | \mathbf{y}_t)}$$

Joint entropies (and distributions) can have high dimension!

Entropy estimates from nearest neighbors [Kraskov et al, 2004]



Nonlinear causality measures (direct and indirect)

Transfer Entropy (TE) [Schreiber, 2000]

Measure the effect of X on Y at one time step ahead, accounting (conditioning) for the effect from its own current state

$$TE_{X \to Y} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t)$$

$$= H(\mathbf{x}_t, \mathbf{y}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) + H(y_{t+1}, \mathbf{y}_t) - H(\mathbf{y}_t)$$

$$= \sum p(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) \log \frac{p(y_{t+1} | \mathbf{x}_t, \mathbf{y}_t)}{p(y_{t+1} | \mathbf{y}_t)}$$

Joint entropies (and distributions) can have high dimension!

Entropy estimates from nearest neighbors [Kraskov et al, 2004]

TE is equivalent to GCI when the stochastic process of (X, Y) is Gaussian [Barnett et al, PRE 2009]



Nonlinear causality measures (direct)

```
driving system: X, response system: Y, conditioning on system Z, Z = \{Z_1, Z_2, \ldots, Z_{K-2}\} join all K-2 z-reconstructed vectors: \mathbf{Z}_t = [\mathbf{z}_{1,t}, \ldots, \mathbf{z}_{K-2,t}]
```

Partial Transfer Entropy (PTE) [Vakorin et al, 2009; Papana et al, 2012]

Measure the effect of X on Y at T times ahead, accounting (conditioning) for the effect from its own current state and the current state of the other variables except X.

$$\begin{aligned} & \mathsf{PTE}_{X \to Y|Z} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t, \mathbf{Z}_t) \\ &= H(\mathbf{x}_t, \mathbf{y}_t | \mathbf{Z}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t | \mathbf{Z}_t) + H(y_{t+1}, \mathbf{y}_t | \mathbf{Z}_t) - H(\mathbf{y}_t | \mathbf{Z}_t) \end{aligned}$$

Joint entropies (and distributions) can have very high dimension!



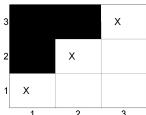
Example: Nonlinear stochastic process

Nonlinear stochastic map:

$$\begin{array}{rcl} x_{1,t} & = & 3.4x_{1,t-1}(1-x_{1,t-1}^2)e^{-x_{1,t-1}^2} + 0.4e_{1,t} \\ x_{2,t} & = & 3.4x_{2,t-1}(1-x_{2,t-1}^2)e^{-x_{2,t-1}^2} + 0.5x_{1,t-1}x_{2,t-1} + 0.4e_{2,t} \\ x_{3,t} & = & 3.4x_{3,t-1}(1-x_{3,t-1}^2)e^{-x_{3,t-1}^2} + 0.3x_{2,t-1}^2 + 0.5x_{1,t-1}^2 + 0.4e_{3,t} \end{array}$$

[Model 7, Gourevich et al, 2006]

true connecntivity matrix



Example: Nonlinear stochastic process

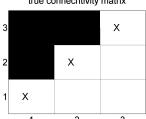
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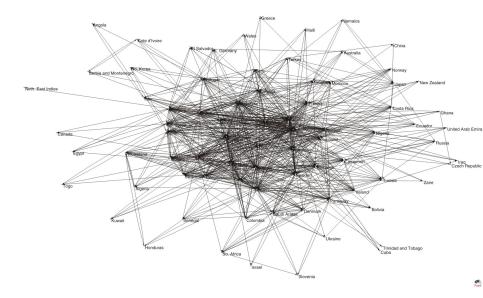
true connecntivity matrix

Estimation of the correct causality effects from the time series?



- Dependence measures in univariate time series
- Interdependence in multivariate time series
- Omplex networks from multivariate time series
- 4 High-dimensional time series: Implications and solutions

Example: Games of world cup 1930 - 2006



Example: Flight connections



Source: https://au.pinterest.com/pin/488077678338752549/

Example: Flight connections



Example: Flight connections

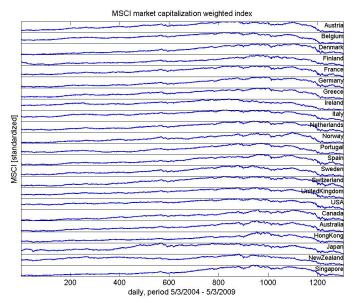


Example: Ship connections



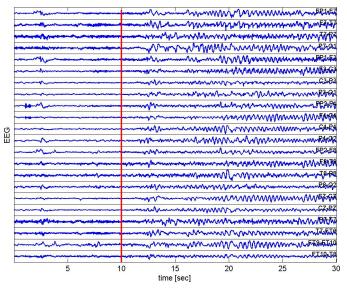
http://i.dailymail.co.uk/i/pix/2014/05/22/article-2636152-1E1A482300000578-442_964x541.jpg

Example: Global financial market





Example: Brain dynamical system

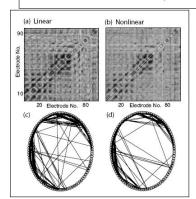


Network: ?

PHYSICAL REVIEW E 79, 061916 (2009)

Network inference with confidence from multivariate time series

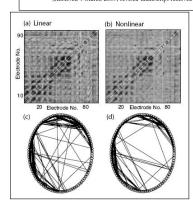
Mark A. Kramer, ^{1,8} Uri T. Eden, ¹ Sydney S. Cash, ² and Eric D. Kolaczyk ¹ Department of Mathematics and Statistics, Boston University, Boston, Massachusetts 02215, USA ² Department of Neurology, Epilepsy Service, Harvard Medical School, ACC 835, Massachusetts General Hospital, 55 Fruit Street, Boston, Massachusetts 02114, USA (Received 9 March 2009; revised manuscript received 14 May 2009; published 11 June 2009)



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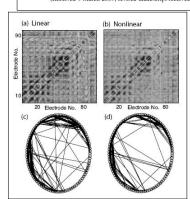


ECoG: "Linear and nonlinear association measures produce similar association matrices and networks."

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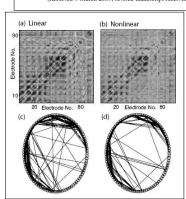


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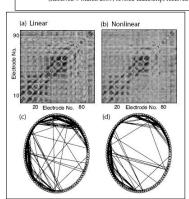
It is important to:

 Use appropriate associate measure ⇒ weighted connection

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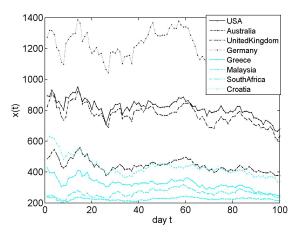
- Use appropriate associate measure ⇒ weighted connection
- Assess the significance of the measure ⇒ binary connection

Example: World financial markets

N=8 world stock markets, daily indices, n=100 days.

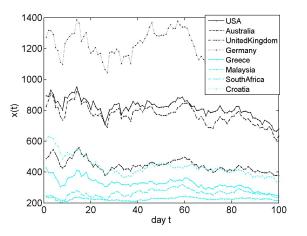
Example: World financial markets

N=8 world stock markets, daily indices, n=100 days.



Example: World financial markets

N=8 world stock markets, daily indices, n=100 days.



Similar indices, links among world stock markets?



Example: World financial markets, correlation coefficient

Upper triangular: sample correlation coefficient r_{ij} .

Lower triangular: p-value for significance test for ρ_{ij} (z-statistic)

	USA	AUS	UK	GER	GRE	MAL	SAF	CRO
USA		0.86	0.92	0.88	0.89	0.33	0.27	0.75
AUS	0		0.91	0.82	0.90	0.56	0.27	0.83
UK	0	0		0.88	0.92	0.40	0.31	0.74
GER	0	0	0		0.84	0.44	0.53	0.61
GRE	0	0	0	0		0.40	0.16	0.82
MAL	0.0008	0	0	0	0		0.54	0.38
SAF	0.0057	0.0065	0.0017	0	0.1154	0		-0.15
CRO	0	0	0	0	0	0.0001	0.1408	

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Almost all indices are strongly correlated.



Example: World financial markets, correlation network

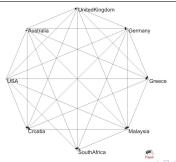
Adjacency matrix, threshold at $\alpha = 0.01$ (multiple testing?)

	USA	AUS	UK	GER	GRE	MAL	SAF	CRO
USA	0	1	1	1	1	1	1	1
AUS	1	0	1	1	1	1	1	1
UK	1	1	0	1	1	1	1	1
GER	1	1	1	0	1	1	1	1
GRE	1	1	1	1	0	1	0	1
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	-					`	•	_
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MAL	1	1	1	1	1	0	1	1
SAF	1	1	1	1	0	1	0	0
CRO	1	1	1	1	1	1	0	0



Example: World financial markets, partial correlation

Upper triangular: partial correlation $r_{ij|K}$, conditioned on all |K| = 6 rest variables.

Lower triangular: p-value for significance test for $\rho_{ij|K}$ (z-statistic)

	USA	AUS	UK	GER	GRE	MAL	SAF	CRO
USA		0.01	0.37	0.27	0.07	-0.27	0.11	0.27
AUS	0.9378		0.42	-0.02	0.15	0.30	0.10	0.38
UK	0.0002	0		0.08	0.36	-0.16	0.08	-0.11
GER	0.0081	0.8469	0.4693		0.38	-0.31	0.66	0.26
GRE	0.4946	0.1392	0.0003	0.0001		0.19	-0.36	0.01
MAL	0.0083	0.0033	0.1232	0.0026	0.0710		0.68	0.46
SAF	0.2908	0.3554	0.4321	0	0.0003	0		-0.70
CRO	0.0079	0.0002	0.3149	0.0099	0.9083	0	0	

Example: World financial markets, partial correlation

Upper triangular: partial correlation $r_{ij|K}$, conditioned on all |K| = 6 rest variables.

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SAF	0.2908	0.3554	0.4321	0	0.0003	0		-0.70
CRO	0.0079	0.0002	0.3149	0.0099	0.9083	0	0	

Correlation between any two indices decreased when conditioned on all others.



Example: Financial markets, partial correlation network

Adjacency matrix, threshold at $\alpha = 0.01$

	USA	AUS	UK	GER	GRE	MAL	SAF	CRO
USA	0	0	1	1	0	1	0	1
AUS	0	0	1	0	0	1	0	1
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GER	1	0	1	0	1	1	1	1
GRE	0	0	1	1	0	0	1	0
MAL	1	1	1	1	0	0	1	1
SAF	0	0	1	1	1	1	0	1
CRO	1	1	1	1	0	1	1	0

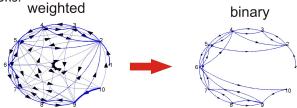
Example: Financial markets, partial correlation network

Adjacency matrix, threshold at $\alpha = 0.01$

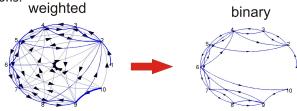
	LICA	ALIC	LUZ	CED	CDE	N 4 A 1	CAE	CDO
	USA	AUS	UK	GER	GRE	MAL	SAF	CRO
USA	0	0	1	1	0	1	0	1
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UK	1	1	0	0	1	0	0	0
GER	1	0	1	0	1	1	1	1
GRE	0	0	1	1	0	0	1	0
MAL	1	1	1	1	0	0	1	1
SAF	0	0	1	1	1	1	0	1
CRO	1	1	1	1	0	1	1	0



Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:

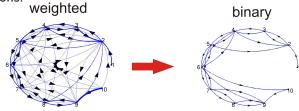


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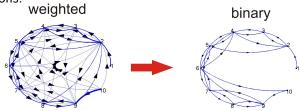
① Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$

Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:



- **①** Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$
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Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:



- **1** Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$
- ② Threshold on the network density, only the d% largest $q(i \rightarrow j)$.
- 3 Significance test on each $q(i \rightarrow j)$. Threshold, e.g. $\alpha = 0.05$ on the p-value of the test.
 - Parametric or resampling test (resampling test for a nonlinear causality measure).

Significance resampling test on $q(i \rightarrow j)$ for each pair (X_i, X_j) .

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K(K-1) significance tests \Longrightarrow Correction for multiple testing

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K(K-1) significance tests \Longrightarrow Correction for multiple testing

Popular choice:

False Discovery Rate (FDR) [Benjamini & Hochberg, 1995]

- ullet K(K-1) p-values in ascending order: $p_{(1)}, p_{(2)}, \ldots, p_{(K(K-1))}$
- Rejection for the k tests with $p \le p_{(k)}$, where $p_{(k)}$ is the largest p-value for which $p_{(k)} < k\alpha/(K(K-1))$.

Significance resampling test on $q(i \rightarrow j)$ for each pair (X_i, X_j) .

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Small p-value can only be obtained with large number of surrogates When K gets large, FDR requires **huge** M (impractical).



Example: coupled Henon maps

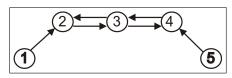
$$x_{1,t+1} = 1.4 - x_{1,t}^2 + 0.3x_{1,t-1}$$

$$x_{i,t+1} = 1.4 - (0.5C(x_{i-1,t} + x_{i+1,t}) + (1 - C)x_{i,t})^2 + 0.3x_{i,t-1}$$

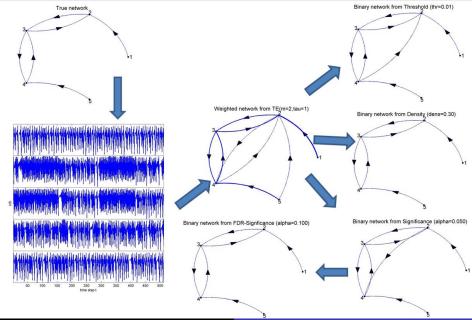
$$x_{K,t+1} = 1.4 - x_{K,t}^2 + 0.3x_{K,t-1}$$

C: coupling strength [Politi & Torcini, 1992]

Network structure for K = 5

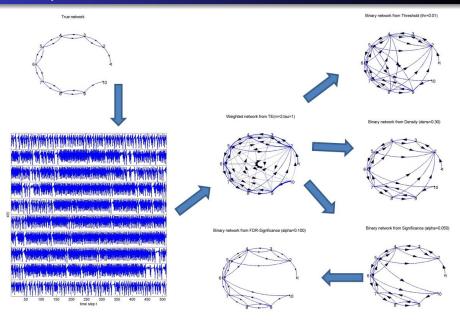


Example, TE, K = 5



- Dependence measures in univariate time series
- 2 Interdependence in multivariate time series
- 3 Complex networks from multivariate time series
- 4 High-dimensional time series: Implications and solutions

Example, TE, K = 10



What if there are many observed variables?

The curse of dimensionality:

• For FDR, in general $M \sim K(K-1)/\alpha$. When K gets large, huge M may be required (impractical).

What if there are many observed variables?

The curse of dimensionality:

- For FDR, in general $M \sim K(K-1)/\alpha$. When K gets large, huge M may be required (impractical).
- For K > 2, bivariate measures are likely to produce false couplings (indirect connections).

Example, TE, K = 20

True network



en (majalematemperiamis)||Italematemperiamis||Italematemperiamis|

Weighted network from TE(m=2,tau=1)

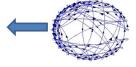




Binary network from Threshold (thr=0.04)

Binary network from Density (dens=0.10)





Binary network from Significance (alpha=0.05

linary network from FDR-Signficance (alpha=0.0!

What if there are many observed variables?

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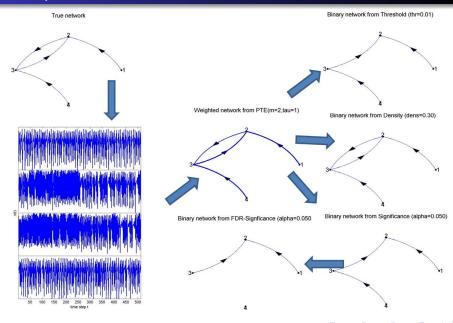
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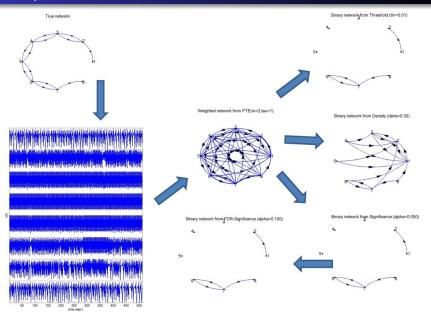
The curse of dimensionality:

- For FDR, in general $M \sim K(K-1)/\alpha$. When K gets large, huge M may be required (impractical).
- For K > 2, bivariate measures are likely to produce false couplings (indirect connections).
- Multivariate measures require long time series, e.g. $\mathsf{PTE}_{X \to Y|Z} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t, \mathbf{Z}_t) \text{ requires the estimation of entropy of } [y_{t+1}, \mathbf{x}_t, \mathbf{y}_t, \mathbf{Z}_t]' \text{ of dimension } 1 + Km.$

Example, PTE, K = 4



Example, PTE, K = 8



Interdependence using Dimension Reduction

K time series $\{x_t, y_t\}_{t=1}^n$ and $\{\mathbf{z}_t\}_{t=1}^n = \{z_{1,t}, z_{2,t}, \dots, z_{K-2,t}\}_{t=1}^n$ driving system: X, response system: Y, conditioning on system Z, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$

If K large with respect to n multivariate Granger causality is problematic ("the curse of dimensionality").

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In CGCI, the VAR model

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + \sum_{i=1}^p A_i \mathbf{z}_{t-i} + e_{U,t}$$

has Kp lagged variables.



Interdependence using Dimension Reduction

K time series $\{x_t, y_t\}_{t=1}^n$ and $\{\mathbf{z}_t\}_{t=1}^n = \{z_{1,t}, z_{2,t}, \dots, z_{K-2,t}\}_{t=1}^n$ driving system: X, response system: Y, conditioning on system Z, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$

If K large with respect to n multivariate Granger causality is problematic ("the curse of dimensionality").

In CGCI, the VAR model

$$y_t = \sum_{i=1}^p a_i y_{t-i} + \sum_{i=1}^p b_i x_{t-i} + \sum_{i=1}^p A_i \mathbf{z}_{t-i} + e_{U,t}$$

has Kp lagged variables.

Statistics methodology: dimension reduction, sparse regression, restricted regression, and sparse/retricted VAR models.



Partial Mutual Information from Mixed Embedding (PMIME) applies dimension reduction using mutual information. The idea [Vlachos & Kugiumtzis, 2010]:

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Similar approaches based on this idea: [Faes et al, PRE 2011; Stamaglia et al, PRE 2012; Runge et al, PRL 2012; Wibral et al, PLOSOne 2013; Runge et al, PRE 2015; edited book of Wibral, Vicente and Lizier "Directed information measures in Neuroscience", Springer, 2014.]

PMIME - 2

The mixed embedding scheme

• Start with an empty embedding vector \mathbf{w}_t^0 , future vector of Y, y_{t+1} , and maximum lag L (or L_x for X, L_y for Y etc) $\mathbf{W}_t = \{x_t, \dots, x_{t-L-1}, y_t, \dots, y_{t-L-1}, z_{1,t}, \dots, z_{K-2,t-L-1}\}$

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• Progressive vector building stops at step j ($\mathbf{w}_t = \mathbf{w}_t^{j-1}$): Criterion of hard threshold:

$$I(y_{t+1}; \mathbf{w}_t^{j-1})/I(y_{t+1}; \mathbf{w}_t^{j}) > A \text{ (here } A = 0.95)$$

Criterion of adaptive threshold: randomization significance test on $I(y_{t+1}; w_t^j | \mathbf{w}_t^{j-1})$

The non-uniform mixed embedding vector of lags of all X, Y, Z for explaining y_{t+1} :

$$\mathbf{w}_{t} = (\underbrace{x_{t-\tau_{x1}}, \dots, x_{t-\tau_{xm_{x}}}}_{\mathbf{w}_{t}^{x}}, \underbrace{y_{t-\tau_{y1}}, \dots, y_{t-\tau_{ym_{y}}}}_{\mathbf{w}_{t}^{y}}, \underbrace{z_{t-\tau_{z1}}, \dots, z_{t-\tau_{zm_{z}}}}_{\mathbf{w}_{t}^{z}})$$

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The causality measure PMIME

$$R_{X \to Y|Z} = \frac{I(y_{t+1}; \mathbf{w}_t^{\times} \mid \mathbf{w}_t^{Y}, \mathbf{w}_t^{Z})}{I(y_{t+1}; \mathbf{w}_t)}$$

 R_{X→Y|Z}: information on the future of Y explained only by X-components of the embedding vector (given the components of Y and Z), normalized with the mutual information of the future of Y and the embedding vector.

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- R_{X→Y|Z}: information on the future of Y explained only by X-components of the embedding vector (given the components of Y and Z), normalized with the mutual information of the future of Y and the embedding vector.
- If \mathbf{w}_t contains no components from X, then $R_{X\to Y|Z}=0$ and X has no direct effect on the future of Y.



Three main advantages of PMIME

• $R_{X \to Y|Z} = 0$ when no significant causality is present, and $R_{X \to Y|Z} > 0$ when it is present [no significance test, no issues with multiple testing!]

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- ⇒ good candidate for causality analysis with many variables

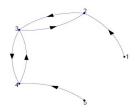


Example: coupled Mackey-Glass

Coupled identical Mackey-Glass delayed differential equations

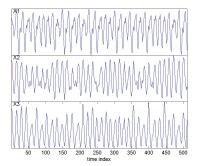
$$\dot{x}_i(t) = -0.1x_i(t) + \sum_{i=1}^K \frac{C_{ij}x_j(t-\Delta)}{1+x_j(t-\Delta)^{10}}$$
 for $i=1,\ldots,K$

$$K = 5$$



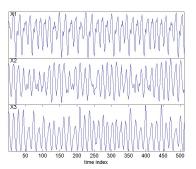
Mackey-Glass, C = 0.2

 $\Delta = 20$

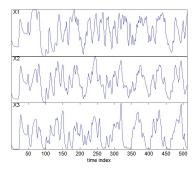


Mackey-Glass, C = 0.2

$$\Delta = 20$$



$$\Delta = 100$$



Mackey-Glass: true/estimated network [Kugiumtzis & Kimiskidis, 2015]

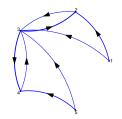
$$K=5$$

True

from PMIME ($\Delta = 20$)

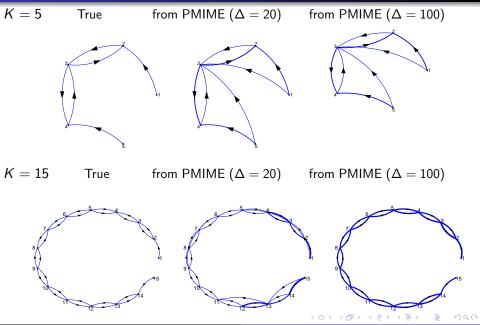
from PMIME ($\Delta=100$)





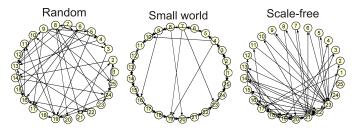


Mackey-Glass: true/estimated network [Kugiumtzis & Kimiskidis, 2015]



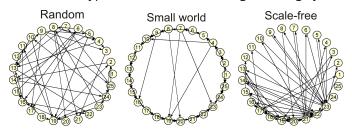
Can different network structures be detected?

Simulation: three types of networks for the generating system



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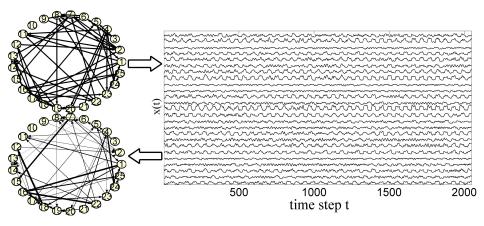


Generating system: coupled Mackey-Glass system, $K=25,~\Delta=100,~C=0.2$ with coupling structure defined by the network type

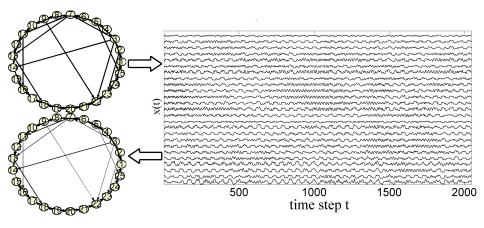
Causality measure: PMIME



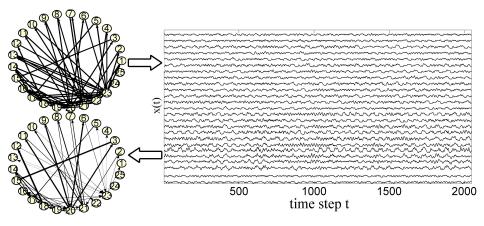
Estimation of the Random Network



Estimation of the Small-World Network



Estimation of the Scale-Free Network



Simulation example:

The network structure undergoes structural change at specific time points:

 $Random \Rightarrow Small-World \Rightarrow Scale-Free$

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Estimation of network characteristics on the PMIME networks

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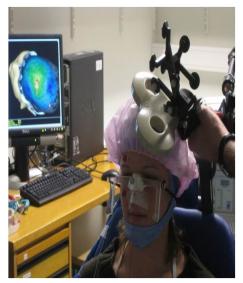
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Estimation of network characteristics on the PMIME networks

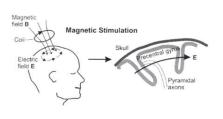
Structural change detection, [Slow], [Middle], [Fast], [Very fast]



EEG and Transcranial Magnetic Stimulation (TMS)



Jointly with Vassilis Kimiskidis, Medical School, AUTh



How does TMS act on epileptic brain connectivity?

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 $\{X_1, X_2, \dots, X_K\}$: K EEG channels, each represents a (sub)system

Practical problems to overcome:

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Application on small time windows ⇒ limited data size

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PMIME addresses all these problems!



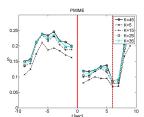
Example: compare PMIME to other measures on EEG

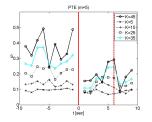
[Kugiumtzis, PRE, 2013]

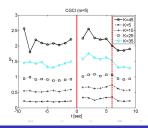
One epileptiform discharge (ED) episode terminated by transcranial magnetic stimulation (TMS), totally 45 channels

- Select randomly a subset of channels.
- Compute the connectivity measures on the subset at each sliding window
- Compute average connectivity strength at each sliding window.
- Repeat the steps above a number of times (here 12).

... for subsets of 5, 15, 25, 35 and once for 45 channels.





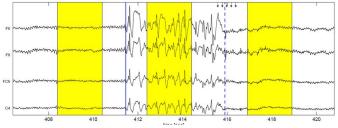


Epileptiform discharges induced by TMS

Preprocessing:

[One Episode]

- replacement of TMS artifact, high order FIR
- rejection of channels with artifacts
- reference to infinity (REST) [Qin et al, ClinNeuroph 2010]
- overlapping windows of 2s, a sliding step of 0.5s

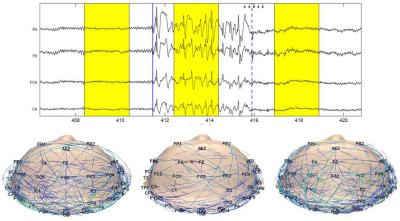


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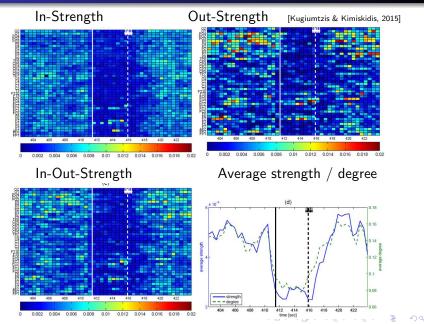
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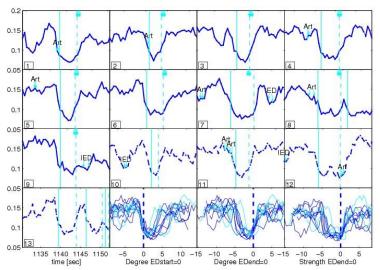
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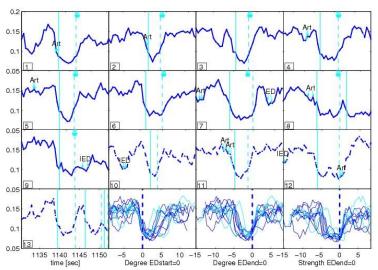
subject 1 with focal seizure, ED episode ends with TMS



Subject 1 with focal seizure, 13 episodes, average degree

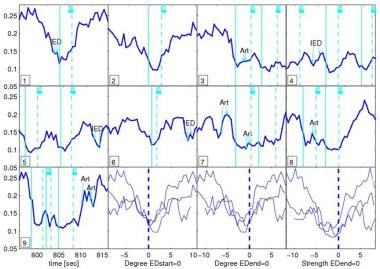


Subject 1 with focal seizure, 13 episodes, average degree



TMS terminates the ED prematurely and restores the network structure as if it would have terminated spontaneously

Subject 2 with focal seizure, 9 episodes, average degree



TMS terminates the ED prematurely and restores the network structure as if it would have terminated spontaneously

Many network indices (totally 78) computed on the PMIMF-causality networks

	THIVIE Causanty Hetworks
Symbol	Description
deg ^m	degree distribution, m=mean,std,skewness,kurtosis
strm	strength distribution, m=mean,std,skewness,kurtosis
TrR_k	transitivity ratio, k=binary undirected (bu),binary directed (bd)
	weighted directed (wd)
$EigC^m$	eigenvector centrality distribution, m=mean,std
λ_k	characteristic path length, k=bd,wd
GE_k	global efficiency, k=bd,wd
ϵ_k^m	eccentricity distribution, m=mean,std and k=bd,wd
rad_k	radius, k=bd,wd
d_k	diameter, k=bd,wd
C_k^m	clustering coefficient distribution,m=mean,std and k=bd,wd
g_k^m	betweenness centrality distribution,m=mean,std and k=bd,wd
$e - g_k^m$	edge betweenness centrality distribution,m=mean,std and k=bd,wd
LE_k^m	local efficiency distribution,m=mean,std and k=bd,wd
3motif(i)	i^{th} motif of 3 nodes, $i=1,2,13$
modul(i)	modularity for i modules, i=2,3,5
$r_{deg}(i, j)$	assortativity coefficient in terms of the degree, i=in,out and j=in,out or i,j=und
$r_{str}(i,j)$	assortativity coefficient in terms of the strength, i=in,out and j=in,out or i,j=und
p_{top}	Rent exponent:topological
p_{ph}	Rent exponent:physical
p_{ee}	Rent exponent:efficient embedding
SW_k	small-worldness, k=bd,wd
kcs	k-core size, k=90-percentile of degree distribution
SCS	s-core size, k=90-percentile of strength distribution
ϕ_k	Rich club coefficient, k=bd,wd
cycprob ₁	fraction of 3-cycles out of 3-paths
$cycprob_2$	probability: non-cyclic 2-path extend to 3-cycle

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For both subjects and pairs preED - ED ED - postED

Measure	AUROC
deg ^{mean}	0.9296
d_{bd}	0.9268
3motif(1)	0.9231
λ_{bd}	0.9231
str ^{mean}	0.9207
3motif(3)	0.9206
3motif(5)	0.9199
LE_{bd}^{mean}	0.9173
GE_{wd}	0.9163
TrR_{wd}	0.9150

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kcs	k-core size, k=90-percentile of degree distribution		
SCS	s-core size, k=90-percentile of strength distribution		
ϕ_k	Rich club coefficient, k=bd,wd		
$cycprob_1$	fraction of 3-cycles out of 3-paths		
$cycprob_2$	probability: non-cyclic 2-path extend to 3-cycle		

For both subjects and pairs preED - ED ED - postED

Measure	AUROC
deg ^{mean}	0.9296
d_{bd}	0.9268
3motif(1)	0.9231
λ_{bd}	0.9231
str ^{mean}	0.9207
3motif(3)	0.9206
3motif(5)	0.9199
LE_{bd}^{mean}	0.9173
GE_{wd}	0.9163
TrR_{wd}	0.9150

subject 1 with genetic generalized epilepsy (GGE)

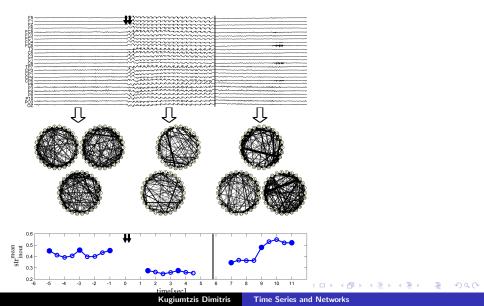
ED induced by TMS, [PMIME on 2s windows]

[Kugiumtzis et al, 2016]

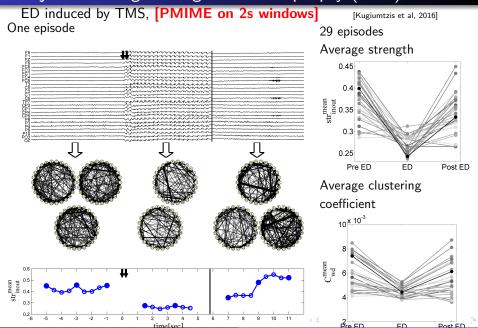
subject 1 with genetic generalized epilepsy (GGE)

ED induced by TMS, [PMIME on 2s windows]
One episode

[Kugiumtzis et al, 2016]



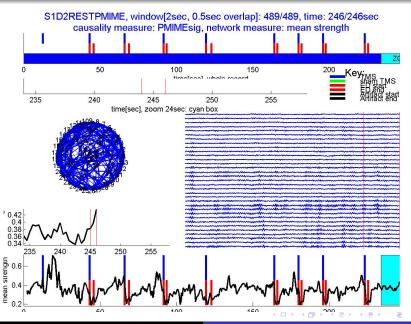
subject 1 with genetic generalized epilepsy (GGE)



Time Series and Networks

Kugiumtzis Dimitris

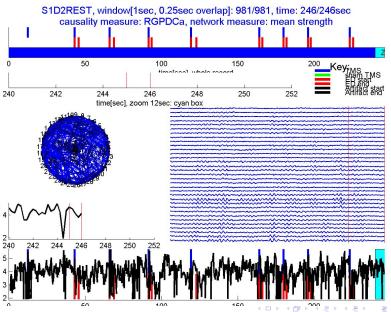
PMIME on 2s window



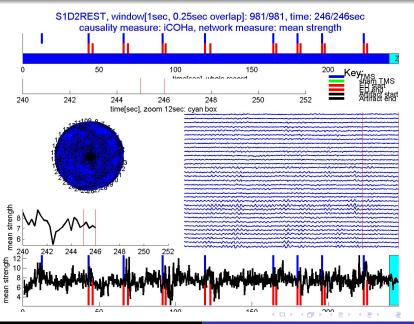
$\mathsf{RGPDC}(\alpha)$ on 1s window

mean strength

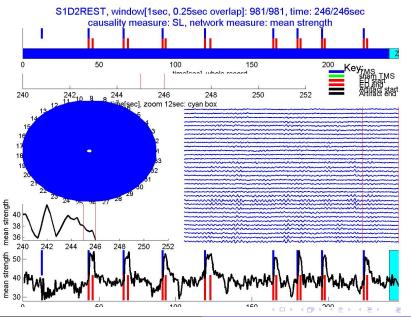
mean strength



$\mathsf{iCoh}(\alpha)$ on 1s window

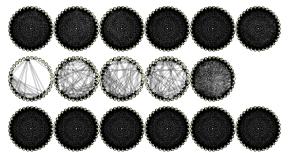


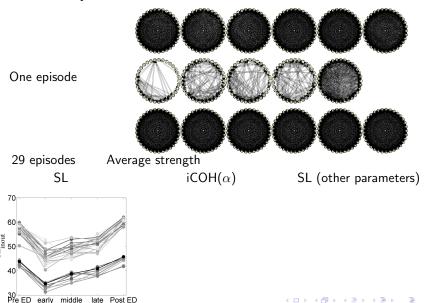
Synchronization Likelihood (SL) on 1s window

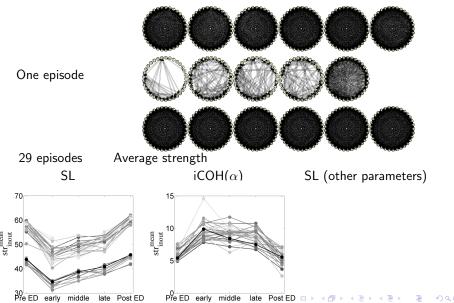


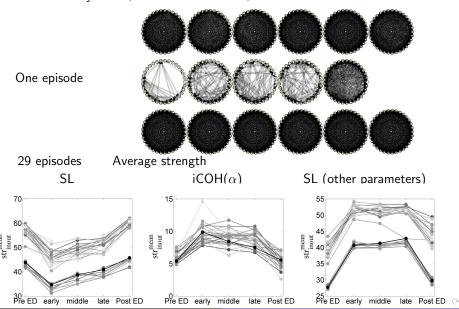
ED induced by TMS, SL on 1s windows,

One episode









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- PMIME is model-free and almost parameter-free, can estimate nonlinear direct causal effects in the presence of many variables
- How can we learn the underlying dynamics of high-dimensional time series?

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Maria Papapetrou, PhD candidate

